



Demand Forecast Update Jemena Gas Networks (NSW) Ltd Access Arrangement - 2020 to 2025

January 2020

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1. Glossary

AER	Australian Energy Regulator
CORE	Core Energy & Resources Pty Ltd
GJ	Gigajoule
HIA	Housing Industry Association
JGN	Jemena Gas Networks (NSW)
NGL	National Gas Law
NGO	National Gas Objective
NGR	National Gas Rules
NSW	New South Wales
Original CORE Submission	The submission in May 2019 based on data to end 2018 (report and model)
Revised CORE submission	This report and related model, updated for 2018-19 data.

2. Introduction

This document has been prepared in response to the AER's Draft Decision (Attachment 12), in relation to the demand forecasts incorporated in the JGN (NSW) 2020-2025 Access Arrangement.

CORE has been engaged by JGN to review the AER's draft decision, to provide an independent response to each issue raised, based on information provided by JGN and sourced from third parties and prepare a revised forecast.

The structure of this document is as follows:

2. Summary and Revised Forecasts

3. Tariff V

3.1 JGN Data Update

3.2 HIA Data Update

3.3 Billing Data Clarification

3.4 HIA One Year Lag

3.5 Suspended Connections (formerly Zero Consuming Meters)

3.6 Volume Boundary Meter Strategy

Attachment 1: Updated Model

Attachment 2: HIA Lag – Statistical Analysis

3. Summary and Revised Forecasts

3.1. Methodology

CORE notes the following statement included in 12.1 of the Draft Decision:

“We are satisfied that the overall demand forecasting methodology applied by JGN’s consultant, CORE Energy & Resources (CORE), is consistent with rule 74(2) of the National Gas Rules (NGR)”.

CORE confirms that it has adopted the same methodology in developing the revised forecasts which are included with this submission as it used in its Original Submission. More specifically the changes CORE has made to the model are:

- updated weather normalisation model based on new 2019-year weather data
- updated the demand forecasting model for:
 - > latest connections and demand data sourced from JGN
 - > latest HIA data

3.2. Tariff V

CORE has noted the following comments by the AER in its Draft Decision:

12.1.1 Additional information sought

“We would like JGN in its revised proposal to:

- *incorporate updated demand and customer forecast based on 2018–19 actual data, and the latest HIA data*
- *provide clarifications on the basis in which the billing data is derived including the source data and associated workings in a separate worksheet*
- *provide further clarifications on the accuracy of using a one-year lag between HIA data and JGN connection, including the likely error margins to the penetration rates*
- *describe the concept of zero consuming meters, its application to the demand and customer forecast, how it differs from disconnections, its relationship with other areas in JGN’s proposal, and the reasons it complies with the National Gas Objective (NGO)*
- *confirm the impact to the demand and customer forecast with and without JGN’s volume boundary meter strategy in which individual hot water meters remain on offer for buildings with centralised hot water systems.*

We recommend that JGN seek independent assurance on any updates to the demand and customer forecast as well as validating that the key inputs used in the demand and customer forecast are fit for purpose”.

CORE confirms that each of the above issues has been addressed and provides independent assurance regarding updates to demand and customer forecast, as well as validating that the key inputs in the demand and customer forecast are fit for purpose.

Each of the matters set out above are addressed individually below.

3.3. Revised Demand Forecast

The following is a summary of the revised forecasts for Tariff V – Residential and Commercial - connections and demand.

The most significant movements between the Revised Submission forecasts and those which formed part of the Original Submission are set out below. CORE notes that the major factor contributing to increased demand for both residential and small business customers is a forecast increase in consumption per connection.

Residential

- Increase in total annual demand by 2025 from 27.55 PJ's to 27.87 PJ's
- Negligible movement in new connections
- Increase in demand per connection by 2025 from 18.3 to 18.5 GJ per customer
- Changes to address the timing of suspended connection in 2019 vs 2018 and changes to reflect change in new connection customer mix as summarised in the tables below.

Small Business

- Increase in total annual demand by 2025 from 13.29 PJ's to 13.59 PJ's
- Minor movement in new connections from 39,630 to 39,689
- Increase in demand per connection by 2025 from 335.3 to 342.4 GJ per customer

The following figures and tables provide a summary of forecast aggregate demand, connections and demand per connection for both residential and small business customers.

Figure 3.1 Residential Connections (1,000's)

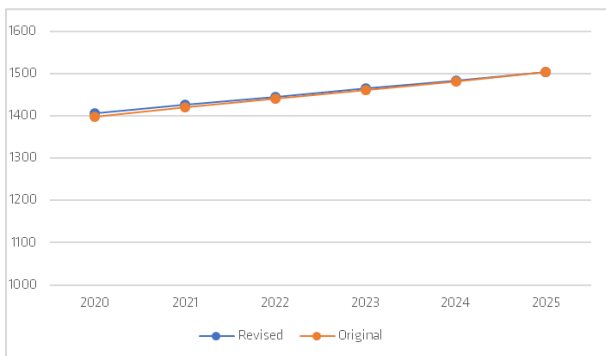


Figure 3.2 Residential Consumption per Connection (G)

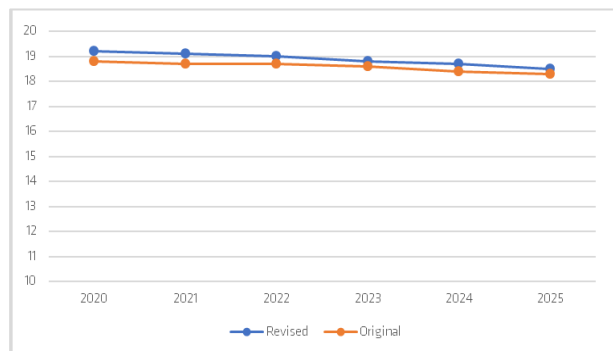


Figure 3.3 Small Business Connections

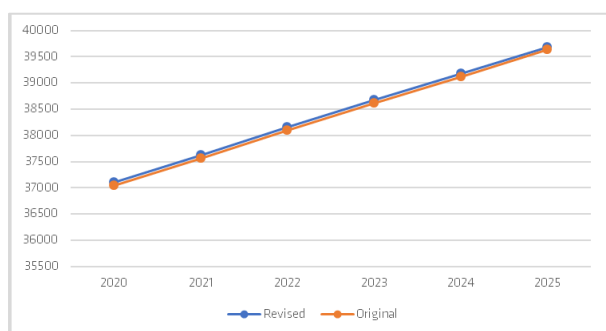
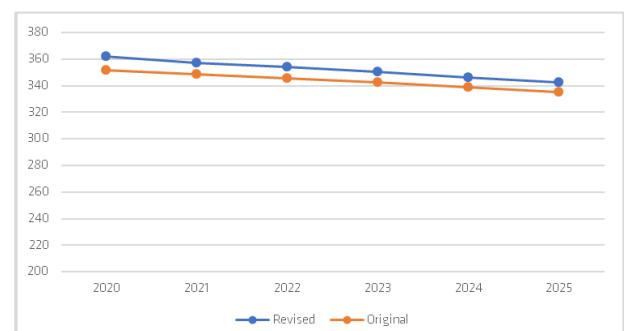


Figure 3.4 Small Business Consumption per Connection (G)



3.4. Revised Submission Forecasts

3.4.1. Residential

Figure 3.5 Residential Demand

Residential	2020	2021	2022	2023	2024	2025
Existing 2018 Connections	26,822,203	26,520,419	26,212,467	25,873,436	25,541,345	25,203,377
New Dwelling Connections Residential Estate	116,542	450,020	768,563	1,062,769	1,349,942	1,635,449
New Dwelling Connections Residential E2G	20,647	82,634	149,150	213,180	274,956	334,417
New Dwelling Connections Residential HRVB	22,450	90,292	162,977	233,581	303,543	374,741
New Dwelling Connections Residential MD	12,166	47,536	82,833	116,539	149,681	182,972
New Dwelling Connections Residential HRVI	16,958	58,390	83,497	104,576	125,471	146,172
Total Demand	27,010,966	27,249,292	27,459,486	27,604,083	27,744,939	27,877,128

Figure 3.6 Residential Connections

Residential	2020	2021	2022	2023	2024	2025
Opening Connections	1,398,848	1,405,425	1,426,103	1,445,411	1,464,509	1,483,589
Disconnections	2,959	3,030	3,084	3,136	3,188	3,240
Disconnections Zero Consuming Connections	27,023	5,162	5,255	5,343	5,431	5,520
Existing 2018 Connections	1,368,867	1,360,675	1,352,336	1,343,856	1,335,237	1,326,477
Net Dwelling Connections Estate	21,154	16,879	15,705	15,632	15,620	16,369
Net Dwelling Connections E2G	3,704	3,697	3,692	3,686	3,681	3,676
Net Dwelling Connections HRVB	160	154	151	150	153	160
Net Dwelling Connections MD	3,787	3,272	3,216	3,192	3,241	3,377
Net Dwelling Connections HRVI	4,795	1,837	1,799	1,782	1,816	1,750
Total New Dwelling Connections	36,559	28,869	27,647	27,578	27,699	28,573
Cumulative New Connections Residential Estate	23,817	43,423	61,904	80,358	98,847	118,132
Cumulative New Connections Residential E2G	4,000	8,000	12,000	16,000	20,000	24,000
Cumulative New Connections Residential HRVB	160	314	466	615	768	928
Cumulative New Connections Residential MD	3,787	7,058	10,275	13,467	16,708	20,085
Cumulative New Connections Residential HRVI	4,795	6,632	8,431	10,213	12,029	13,779
Total Connections	1,405,425	1,426,103	1,445,411	1,464,509	1,483,589	1,503,402
% Growth	0.47%	1.47%	1.35%	1.32%	1.30%	1.34%
Net Connections	6,577	20,678	19,307	19,098	19,080	19,813
Total Disconnections (Includes ZCM's)	29,981	8,191	8,339	8,480	8,619	8,760

Figure 3.7 Residential Consumption per Connection

Residential	2020	2021	2022	2023	2024	2025
Existing 2018	19.6	19.5	19.4	19.3	19.1	19.0
New Dwelling Connections Residential Estate	4.9	10.4	12.4	13.2	13.7	13.8
New Dwelling Connections Residential E2G	5.2	10.3	12.4	13.3	13.7	13.9
New Dwelling Connections Residential HRVB	140.2	287.1	350.1	379.7	395.4	403.8
New Dwelling Connections Residential MD	3.2	6.7	8.1	8.7	9.0	9.1
New Dwelling Connections Residential HRVI	3.5	8.8	9.9	10.2	10.4	10.6
Weighted Average Demand per Connection	19.2	19.1	19.0	18.8	18.7	18.5
% Growth	1.97%	-0.58%	-0.57%	-0.78%	-0.78%	-0.85%

3.4.2. Small Business

Figure 3.8 Small Business Demand

Small Business	2020	2021	2022	2023	2024	2025
Existing 2018	13,379,748	13,177,837	13,005,214	12,804,611	12,566,545	12,338,121
New Small Business	56,868	254,273	500,373	749,547	999,393	1,252,190
Total Demand	13,436,616	13,432,111	13,505,587	13,554,159	13,565,938	13,590,312
% Growth	-4.08%	-0.03%	0.55%	0.36%	0.09%	0.18%

Figure 3.9 Small Business Connections

Small Business	2020	2021	2022	2023	2024	2025
Opening Connections	36,596	37,104	37,629	38,151	38,668	39,180
Disconnections	320	302	306	311	315	319
Disconnections Zero Consuming Connections						
Existing 2018 Connections	36,276	35,974	35,668	35,357	35,042	34,723
New Small Business Connections	828	828	828	828	828	828
Cumulative New Commercial Connections	828	1,655	2,483	3,310	4,138	4,966
Total Connections	37,104	37,629	38,151	38,668	39,180	39,689
% Growth	1.39%	1.42%	1.39%	1.36%	1.33%	1.30%
Net Connections	508	526	521	517	513	509
% Growth	-11.55%	3.52%	-0.81%	-0.81%	-0.81%	-0.81%

Figure 3.10 Small Business Consumption per Connection

Small Business	2020	2021	2022	2023	2024	2025
Existing 2018	368.8	366.3	364.6	362.2	358.6	355.3
New Small Business	68.7	153.6	201.5	226.4	241.5	252.2
Weighted Average Demand per Connection	362.1	357.0	354.0	350.5	346.2	342.4
% Growth	-5.40%	-1.43%	-0.83%	-0.98%	-1.22%	-1.10%

4. Tariff V

4.1. Data Update

4.1.1. Connection and Demand Data Update

CORE confirms that JGN has provided updated data for the 2019 year, for Tariff V - residential and commercial customers, including both connection and demand data, in order to develop this Revised Submission. Further CORE confirms that the data is consistent in nature with that used to develop forecasts for the Original CORE submission and has been applied in a consistent manner by CORE to derive the Revised Submission.

This update resulted in an increase in residential and small business demand as referred to above.

CORE refers to the Revised Model which is included as Attachment 1.

4.1.2. HIA Data Update

CORE confirms that it has accessed and applied latest available HIA data to derive Revised Forecasts of Tariff V, Residential connections, including allocation between dwelling types (due to variable levels of gas consumption of each dwelling type). Further CORE confirms that it has applied this data in a manner consistent with the approach used to derive forecasts in the Original Submission.

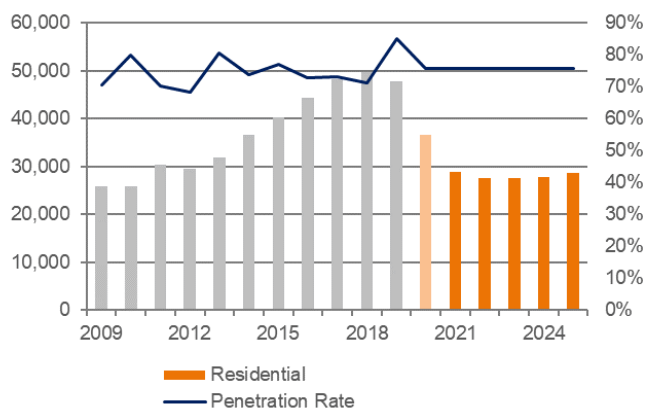
CORE refers to the Revised Model which is included as Attachment 1

4.2. Billing Data and Penetration Rate Explanation

4.2.1. Overview of Approach

- CORE's penetration rate (the proportion of new NSW dwellings which will connect to the JGN network) is based on two sets of data: actual connection data (for new homes, medium density, and individually metered high-rise connections) and estimates of the number of dwellings connected via JGN's volume boundary metering product.
- CORE takes the data and formats it to meet the requirements of CORE's analytical models.
- The final data by customer category is compared against the latest HIA data (by dwelling category) to arrive at the best estimate of Penetration rates for each customer type, across the JGN network
- CORE reiterates that there is a difference between the dwelling connection rate and the connection penetration rate due to the impact of VB connection type which serve multiple dwellings from one connection. The variance between the values, following the 2019 data update, is summarised below:

Figure 4.1 Residential Dwelling Connection Rate vs Connection Penetration Rate



4.2.1.1 High Rise Dwellings

JGN requested that Core provide assurance over the estimates of the number of dwellings and sites connected for each high-rise sub-segment. CORE was not asked to review annual connections as this data has been reported to the AER and audited by KMPG as part of JGN's RIN responses.

The high-rise subsegments are as follows:

- Volume boundary – sites with a single volume boundary meter.
- Instantaneous hot water – sites with individual gas meters for each dwelling.
- Centralised hot water – sites with individual metering for each dwelling (hot water meters or hot water and gas meters).
- Hybrid (volume boundary and centralised hot water) – sites which have both a volume boundary meter and individual hot water meters.

JGN's approach to estimating the number of dwellings and sites is to:

- Use billing data where there was a 1-1 relationship between the number of connections and dwellings or sites. For instance:
 - > Each high-rise VI connection is for a single dwelling.
 - > Each VB connection is for a single site.
- Supplement this data set JGN compiled data from connections applications to identify:
 - > How many dwellings were supplied by each VB connection. JGN calculated this number by adding up the number of dwellings listed in each VB connection application.
 - > How many sites were configured with individual gas or hot water and gas metering. JGN calculated this number by adding up the number of connection applications for these VI connections on the assumption that there is a single connection application for each site.
- Use billing information to identify:
 - > Whether a VI connection was an instantaneous hot water site or a centralised hot water site. This was done by allocating connections based on whether there was a hot water read or not.
 - > Whether a VB connection is a volume boundary or hybrid site based on connection application data. Where a VB connection was requested together with a VI connection which has a hot water read this VB connection was allocated to hybrid.

Table 4.1 Estimation methodology for the high-rise sub-segment

	Dwelling numbers	Number of sites
Volume boundary	Number of dwellings listed in each VB connection application	Number of VB connections where there was no associated VI connection with a hot water read
Instantaneous hot water	Number of high-rise VI connections without a hot water read	Number of connection applications for the associated VI connections
Centralised hot water	Number of high-rise VI connections with a hot water read	Number of connection applications for the associated VI connections
Hybrid (volume boundary and centralised hot water)	Number of high-rise VI connections	Number of VB connections where there was associated VI connection with a hot water read

4.2.2. CORE Conclusion

CORE has undertaken an interview with JGN team members responsible for the relevant data and undertaken a 'walk through' the process and has received relevant data to illustrate the process at a highly granular level.

CORE confirms that it has reviewed the approach to developing the assumed penetration rate, based on billing data and notes that dwelling numbers and connection numbers are consistent and has assessed the approach to be reasonable and consistent with rule 74(2) of the NGR.

The final penetration rates relied upon to develop the Revised Forecasts are included within the Revised Model in Attachment 1.

4.3. Reliance on One Year Lag on HIA Data for Residential Connection Forecasts

4.3.1. Introduction

CORE has considered all reasonable approaches in developing a forecast of dwelling completions, connections and thus penetration rates.

CORE has consistently adopted a one-year lag of HIA data (commencements) to derive an estimate of housing completions (the potential timing of connection) on the following basis:

- The HIA is one of few reputable data providers available in the Australian market. CORE has used other data from a specialist provider in the past and found results less accurate;
- 12 months is a prudent estimate of the time from commencement to completion of an 'average' dwelling (see below);
- This approach has been used and accepted in prior regulatory submissions;

- Reference to ABS data on average dwelling completion times in NSW¹:
 - 7 quarters for new flats, units and apartments
 - 3 quarters for townhouses
 - 2-3 quarters for houses
- There is a statistically significant relationship between JGN connection data and HIA data with a one-year lag - see Attachment 2; and
- CORE is not aware of a suitable alternative which is consistent with rule 74(2) of the NGR.

4.3.2. Use of One Year Lag

Building starts are published by the Housing Industry Association (HIA) on a quarterly basis. Each release details the number of residential dwellings commenced during any given quarter, further detailing the housing type each dwelling commencement falls under. Core has established a lag period of 1 year between the listing of a dwelling commencement and the date at which a dwelling is first billed. This lag period has been determined based on the amount of time taken for a contractor to register a connection request, the installation of the inlet service by a sub-contractor followed by the request and subsequent installation of a gas meter.

Core has determined that the construction time of a new build is the most limiting factor when considering the time at which a gas meter may be installed.

Core recognises that each new dwelling commencement will vary in how long it takes to connect. With multistorey houses and apartment blocks taking considerably longer than a single-story build. Core has therefore investigated the complete range of timeframes quoted for new dwellings to be complete, occupiers moved in and consumption to be first billed. Without considering the pre-construction and planning stages of a new build, approximately 8 months of construction is estimated for single story houses in Australia^{2 3 4} whilst estimates closer to 17months have been quoted for multistorey projects⁵. Research has suggested this 8-month low estimate has been rising in Australia from around 7 months during 2016⁶. Furthermore, the Australian government whilst not giving an estimated timeline quotes that, 'very few new builds fit within their timeliness' because of increasing council legislation, contractor and weather-related delays⁷.

Considering the HIA data is annual, and the significant length of time taken for a housing start to reach a stage of development ready for a occupier to move in Core considers the aforementioned 1 year time lag to be the best estimate of the average time it takes to be installed following commencement and is considered a prudent assumption.

¹<https://www.abs.gov.au/ausstats/abs@.nsf/Latestproducts/8752.0Feature%20Article3Jun%202019?opendocument&tabname=Summary&prodno=8752.0&issue=Jun%202019&num=&view=>

² <https://www.homeloanexperts.com.au/home-loan-articles/stages-of-construction/>

³ <https://www.newhomesguide.com.au/articles/2016/12/20/average-build-time-for-a-house>

⁴ <https://www.novushomes.com.au/blog/how-long-does-it-take-build-new-home-perth>

⁵ <https://www.realestate.com.au/advice/long-take-build-house/>

⁶ <https://www.ruralcoproperty.com.au/2016/04/15/a-typical-timeline-for-how-long-it-takes-to-build-a-new-home/>

⁷ <https://www.yourhome.gov.au/you-begin/construction-process>

4.3.3. Conclusion

CORE considers the use of HIA data and a one-year lag to be a prudent basis for deriving estimates of future residential connections and the estimates included with the Revised Model to be reasonable and consistent with Rule 74(2) of the NGR.

4.4. Suspended Connections (formerly Zero Consuming Meters)

The general description – suspended connections, relates to customers who have been suspended and therefore had their service suspended (physical suspension via a wad being inserted to prevent gas flow to the connection).

4.4.1. JGN Approach

During 2019, JGN and retailers agreed that network charges would cease 20 business days after suspension of a MIRN. This change was implemented in October 2019 and applies to MIRNS which had previously been suspended, and to MIRNS disconnected after that date. When a MIRN is reconnected, network charges resume from the date of reconnection. In preparing the demand forecast for JGN's Draft Proposal the label "zero-consuming MIRNS" or "ZCM" was used to refer to these MIRNS as gas cannot be consumed at the premises because the meter is locked or wadded. This label has been revised to suspended connections to avoid confusion.

In the context of the demand forecast, the following terminology is used:

1. "Suspended Disconnections" are MIRNS which have been suspended (typically through wadding or locking of the meter). Typically, the MIRN is disconnected at the request of the retailer due to non-payment, customer move-out or failure of the customer to set up a retail agreement.

2. "Disconnected MIRNS" are MIRNS which have been permanently disconnected (otherwise referred to as "abolished"). With these MIRNS, the service to the property is permanently removed (e.g. by isolation of the main in the street) and the MIRN is removed from the market.

The proposed 2020-2025 RSA submitted by JGN in June 2019 specifies that the MIRN would be removed from the Customer List and charges will cease 20 Business Days after suspension. Following consultation with retailers, JGN has agreed that the charges for the MIRN will cease from the date of suspension – this change will be reflected in the updated RSA which will be lodged as part of the response to the Draft Decision.

JGN considers that treatment of suspended connections not incurring network charge in this manner is consistent with the NGR because it improves the accuracy of the forecast on which reference tariffs are based.

4.4.2. Modelling approach

It should be noted that the estimation was undertaken by CORE based on information provided by JGN, relating to historical billing records, including analysis of MIRNS which were suspended (temporarily disconnected)..

The removal of MIRNS from network charges is different from the historical situation that applied prior to October 2019, where network charges continued to apply to sites which had been suspended.

CORE notes that the adjustment to 21,982 customers in 2019 related to customers that had already been suspended but the billing of fixed charges stopped in October. Therefore, adjustments were required for connections only and not total consumption.

Core has assessed the historical suspended connection data and identified that approximately 0.36% or 5,000+ customers (who previously used gas) become suspended each year and cease consuming gas. When this occurs in the future JGN will lose the load from the customer and stop billing fixed charges.

By way of illustration, CORE has identified 7,216 connections would have been suspended (for at least 20 days) at a point in time during 2018 financial year and 10,667 during the 2017 financial year. Further analysis indicates that approximately 0.36% of opening connections are expected to be suspended on average during a year.

4.4.3. Conclusion

CORE has undertaken a review of data provided by JGN, calculated the estimates to be included with the Revised Model and determines them to be reasonable and consistent with Rule 74(2) of the NGR.

4.5. Volume Boundary Meter Strategy

4.5.1. Background

The high-rise market can be dissected into the following segments:

- Volume boundary – sites with a single volume boundary meter.
- Instantaneous hot water – sites with individual gas meters for each dwelling.
- Centralised hot water – sites with individual metering for each dwelling (hot water meters or hot water and gas meters).
- Hybrid (volume boundary and centralised hot water) – sites which have both a volume boundary meter and individual hot water meters.

The proportion of high-rise sub-segments can be forecast by analysing the trends observed over the last three years (since JGN's volume boundary metering product was introduced) and by taking into account the factors influencing developer decision making.

Over the last three years there has been a steady rise in the number of dwellings taking up the volume boundary product. The proportion of high-rise dwellings utilising volume boundary has risen by 14% in 2017 to 51% in 2019. There has also been a reduction in the proportion of dwellings being supplied with individual hot water meters (centralised hot water and hybrid sites combined) – from 59% to 41% over the same time period.

The number of high-rise sites with instantaneous hot water systems has also fallen from 26% in 2017 to 9% in 2019. It is thought this may be due to embedders increasingly marketing their product to smaller sites as well as changes to the Australian Standards around the ventilation requirements of gas hot water heaters on balconies.

The gradual increase in the take-up of the volume boundary product is consistent with high-rise construction timelines where projects can take anywhere from 2-5 years to construct⁸. Although the volume boundary product was available from 2016 many

⁸ <https://rba.gov.au/publications/bulletin/2017/jun/pdf/bu-0617-1-houses-and-apartments-in-australia.pdf>

of the developments completed in 2017 and 2018 had already made their design decisions by 2016 and could not take advantage of the volume boundary product.

JGN expects the proportion of dwellings supplied via volume boundary metering to continue to increase as developers communicate and demonstrate that they are motivated by associated reduction in construction costs and space savings, relative to individual hot water metering.

Based on the observed historical trends and the latest market developments Core forecasts that:

- the proportion of Instantaneous hot water dwellings will continue to fall (following a demonstrated historical trend), from 9% to 5% by 2021, after which time the percentage is expected to remain relatively constant;
- the proportion of centralised hot water dwellings will fall to 5% as only a minority of developers will opt to install individual hot water metering;
- the number of hybrid sites will fall to 1%, reflecting the historic ratio of hybrid to centralised hot water sites; and
- the residual number of sites (89% by 2022) will be supplied by volume boundary metering.

4.6. Independent assurance on updates as well as validating key inputs are fit for purpose

CORE has completed an independent assessment of the JGN data used to develop updates of the forecasts of both Tariff V (Residential and Commercial) connection and demand and other input data (including third party).

CORE has undertaken a review of the adjustment to the model by the AER referred to in its Draft Decision and observed that the change in demand is as expected. In this regard CORE notes a number of interrelations between elements of the model.

CORE confirms that the JGN data used is reasonable and the other input data is fit for purpose.

Attachment 1 - Revised Model

Attachment 2 | Statistical Analysis - HIA One Year Lag Relationship with JGN Connections



```

name: <unnamed>
log: S:\Cameron Kirkpatrick\HIA lag\HIA_lags.smcl
log type: smcl
opened on: 10 Dec 2019, 09:32:20
    
```

```

1 . insheet using "S:\Cameron Kirkpatrick\HIA lag\HIA_lags.csv"
   (12 vars, 11 obs)
2 . tabulate year
    
```

year	Freq.	Percent	Cum.
2009	1	9.09	9.09
2010	1	9.09	18.18
2011	1	9.09	27.27
2012	1	9.09	36.36
2013	1	9.09	45.45
2014	1	9.09	54.55
2015	1	9.09	63.64
2016	1	9.09	72.73
2017	1	9.09	81.82
2018	1	9.09	90.91
2019	1	9.09	100.00
Total	11	100.00	

```

3 . gen time = _n
4 . tsset year
   time variable: year, 2009 to 2019
   delta: 1 unit
    
```

```

5 .
6 . *all categories of HIA forecast regressed against JGN recorded connections
7 . reg jgn a
    
```

Source	SS	df	MS	Number of obs =	11
Model	1.4726e+09	1	1.4726e+09	F(1, 9) =	50.46
Residual	262666895	9	29185210.6	Prob > F =	0.0001
Total	1.7353e+09	10	173525941	R-squared =	0.8486
				Adj R-squared =	0.8318
				Root MSE =	5402.3

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
a	1.924004	.2708607	7.10	0.000	1.311274 2.536733
_cons	-2842.548	6268.832	-0.45	0.661	-17023.63 11338.53

```

8 . estat ic
   Akaike's information criterion and Bayesian information criterion
    
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-109.0451	2	222.0902	222.886

Note: N=Obs used in calculating BIC; see [R] BIC note

9 . reg jgn b

Source	SS	df	MS			
Model	457820739	1	457820739	Number of obs = 11		
Residual	1.2774e+09	9	141937630	F(1, 9) = 3.23		
Total	1.7353e+09	10	173525941	Prob > F = 0.1061		
				R-squared = 0.2638		
				Adj R-squared = 0.1820		
				Root MSE = 11914		

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
b	10.56486	5.88254	1.80	0.106	-2.742368	23.87209
_cons	13300.52	15379.34	0.86	0.410	-21489.96	48091.01

10. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-117.7446	2	239.4892	240.2849

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

11. reg jgn c

Source	SS	df	MS			
Model	1.6583e+09	1	1.6583e+09	Number of obs = 11		
Residual	76933899.2	9	8548211.02	F(1, 9) = 194.00		
Total	1.7353e+09	10	173525941	Prob > F = 0.0000		
				R-squared = 0.9557		
				Adj R-squared = 0.9507		
				Root MSE = 2923.7		

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
c	7.776538	.5583275	13.93	0.000	6.513513	9.039562
_cons	9797.053	2351.27	4.17	0.002	4478.111	15115.99

12. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-102.2914	2	208.5828	209.3786

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

13. reg jgn d

Source	SS	df	MS			
Model	381490843	1	381490843	Number of obs = 11		
Residual	1.3538e+09	9	150418730	F(1, 9) = 2.54		
Total	1.7353e+09	10	173525941	Prob > F = 0.1457		
				R-squared = 0.2198		
				Adj R-squared = 0.1332		
				Root MSE = 12265		

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
d	-7.262075	4.56005	-1.59	0.146	-17.57762	3.053475
_cons	50153.94	7285.37	6.88	0.000	33673.29	66634.59

14. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-118.0638	2	240.1275	240.9233

Note: N=Obs used in calculating BIC; see **[R] BIC note**

15. reg jgn e

Source	SS	df	MS	Number of obs =	11
Model	305460712	1	305460712	F(1, 9) =	1.92
Residual	1.4298e+09	9	158866522	Prob > F =	0.1989
				R-squared =	0.1760
				Adj R-squared =	0.0845
Total	1.7353e+09	10	173525941	Root MSE =	12604

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
e	19.74597	14.24023	1.39	0.199	-12.46767 51.95961
_cons	19434.86	15420.11	1.26	0.239	-15447.84 54317.57

16. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-118.3643	2	240.7286	241.5244

Note: N=Obs used in calculating BIC; see **[R] BIC note**

17. reg jgn f

Source	SS	df	MS	Number of obs =	11
Model	1.2373e+09	1	1.2373e+09	F(1, 9) =	22.37
Residual	497921736	9	55324637.4	Prob > F =	0.0011
				R-squared =	0.7131
				Adj R-squared =	0.6812
Total	1.7353e+09	10	173525941	Root MSE =	7438.1

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
f	1.08289	.2289809	4.73	0.001	.5648987 1.60088
_cons	20821.98	4663.224	4.47	0.002	10273.03 31370.92

18. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-112.5626	2	229.1253	229.9211

Note: N=Obs used in calculating BIC; see **[R] BIC note**

19. reg jgn g

Source	SS	df	MS			
Model	1.4684e+09	1	1.4684e+09	Number of obs =	11	
Residual	266820727	9	29646747.4	F(1, 9) =	49.53	
Total	1.7353e+09	10	173525941	Prob > F =	0.0001	
				R-squared =	0.8462	
				Adj R-squared =	0.8292	
				Root MSE =	5444.9	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
g	1.923994	.2733785	7.04	0.000	1.305569	2.542419
_cons	-2907.05	6335.384	-0.46	0.657	-17238.68	11424.59

20. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-109.1314	2	222.2628	223.0586

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

21. reg jgn h

Source	SS	df	MS			
Model	1.2376e+09	1	1.2376e+09	Number of obs =	11	
Residual	497631674	9	55292408.3	F(1, 9) =	22.38	
Total	1.7353e+09	10	173525941	Prob > F =	0.0011	
				R-squared =	0.7132	
				Adj R-squared =	0.6814	
				Root MSE =	7435.9	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
h	.9056124	.1914168	4.73	0.001	.4725975	1.338627
_cons	15141.06	5743.309	2.64	0.027	2148.787	28133.32

22. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-112.5594	2	229.1189	229.9147

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

23. reg jgn total

Source	SS	df	MS			
Model	1.3346e+09	1	1.3346e+09	Number of obs =	11	
Residual	400698283	9	44522031.5	F(1, 9) =	29.98	
Total	1.7353e+09	10	173525941	Prob > F =	0.0004	
				R-squared =	0.7691	
				Adj R-squared =	0.7434	
				Root MSE =	6672.5	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
total	.6266005	.1144482	5.47	0.000	.3677007	.8855003
_cons	8823.294	6066.456	1.45	0.180	-4899.983	22546.57

24. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-111.3679	2	226.7357	227.5315

Note: N=Obs used in calculating BIC; see **[R] BIC note**

25.

26. *all except bde providing a significant relationship

27. *note AIC indicators significantly higher than in later model structures

28.

29. reg jgn a

Source	SS	df	MS			
Model	1.4726e+09	1	1.4726e+09	Number of obs =	11	
Residual	262666895	9	29185210.6	F(1, 9) =	50.46	
Total	1.7353e+09	10	173525941	Prob > F =	0.0001	
				R-squared =	0.8486	
				Adj R-squared =	0.8318	
				Root MSE =	5402.3	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
a	1.924004	.2708607	7.10	0.000	1.311274	2.536733
_cons	-2842.548	6268.832	-0.45	0.661	-17023.63	11338.53

30. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: fitted values of jgn

chi2(1) = 0.94

Prob > chi2 = 0.3329

31. estat imtest, white

White's test for Ho: homoskedasticity

against Ha: unrestricted heteroskedasticity

chi2(2) = 0.80

Prob > chi2 = 0.6711

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	0.80	2	0.6711
Skewness	4.52	1	0.0336
Kurtosis	0.29	1	0.5921
Total	5.60	4	0.2310

32. dwstat

Durbin-Watson d-statistic(2, 11) = 1.357518

33. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	0.020	1	0.8878

H0: no serial correlation

34. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	0.091	2	0.9557

H0: no serial correlation

35. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn

Ho: model has no omitted variables

F(3, 6) = 1.39
 Prob > F = 0.3329

36. estat vif

Variable	VIF	1/VIF
a	1.00	1.000000
Mean VIF	1.00	

37. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-109.0451	2	222.0902	222.886

Note: N=Obs used in calculating BIC; see [R] BIC note

38.

39. reg jgn c

Source	SS	df	MS	Number of obs =	11
Model	1.6583e+09	1	1.6583e+09	F(1, 9) =	194.00
Residual	76933899.2	9	8548211.02	Prob > F =	0.0000
Total	1.7353e+09	10	173525941	R-squared =	0.9557
				Adj R-squared =	0.9507
				Root MSE =	2923.7

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
c	7.776538	.5583275	13.93	0.000	6.513513 9.039562
_cons	9797.053	2351.27	4.17	0.002	4478.111 15115.99

40. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
 Ho: Constant variance
 Variables: fitted values of jgn

chi2(1) = 0.40
 Prob > chi2 = 0.5267

41. estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

chi2(2) = 1.53
 Prob > chi2 = 0.4642

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	1.53	2	0.4642
Skewness	0.38	1	0.5402
Kurtosis	1.74	1	0.1876
Total	3.65	4	0.4560

42. dwstat

Durbin-Watson d-statistic(2, 11) = 2.427626

43. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	1.121	1	0.2897

H0: no serial correlation

44. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	1.123	2	0.5702

H0: no serial correlation

45. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn
 Ho: model has no omitted variables

F(3, 6) = 0.61
 Prob > F = 0.6342

46. estat vif

Variable	VIF	1/VIF
c	1.00	1.000000
Mean VIF	1.00	

47. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-102.2914	2	208.5828	209.3786

Note: N=Obs used in calculating BIC; see **[R] BIC note**

48.

49. reg jgn f

Source	SS	df	MS			
Model	1.2373e+09	1	1.2373e+09	Number of obs =	11	
Residual	497921736	9	55324637.4	F(1, 9) =	22.37	
Total	1.7353e+09	10	173525941	Prob > F =	0.0011	
				R-squared =	0.7131	
				Adj R-squared =	0.6812	
				Root MSE =	7438.1	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
f	1.08289	.2289809	4.73	0.001	.5648987	1.60088
_cons	20821.98	4663.224	4.47	0.002	10273.03	31370.92

50. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance
 Variables: fitted values of jgn

chi2(1) = 1.74
 Prob > chi2 = 0.1869

51. estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

chi2(2) = 2.03
 Prob > chi2 = 0.3618

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	2.03	2	0.3618
Skewness	2.51	1	0.1133
Kurtosis	2.68	1	0.1015
Total	7.22	4	0.1245

52. dwstat

Durbin-Watson d-statistic(2, 11) = .6638264

53. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	7.003	1	0.0081

H0: no serial correlation

54. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	7.123	2	0.0284

H0: no serial correlation

55. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn

Ho: model has no omitted variables

F(3, 6) = 0.43
 Prob > F = 0.7390

56. estat vif

Variable	VIF	1/VIF
f	1.00	1.000000
Mean VIF	1.00	

57. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-112.5626	2	229.1253	229.9211

Note: N=Obs used in calculating BIC; see [R] BIC note

58.

59. reg jgn g

Source	SS	df	MS	Number of obs =	11
Model	1.4684e+09	1	1.4684e+09	F(1, 9) =	49.53
Residual	266820727	9	29646747.4	Prob > F =	0.0001
Total	1.7353e+09	10	173525941	R-squared =	0.8462
				Adj R-squared =	0.8292
				Root MSE =	5444.9

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
g	1.923994	.2733785	7.04	0.000	1.305569 2.542419
_cons	-2907.05	6335.384	-0.46	0.657	-17238.68 11424.59

60. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
 Ho: Constant variance
 Variables: fitted values of jgn

 chi2(1) = 0.92
 Prob > chi2 = 0.3381

61. estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

 chi2(2) = 0.80
 Prob > chi2 = 0.6694

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	0.80	2	0.6694
Skewness	4.44	1	0.0351
Kurtosis	0.20	1	0.6512
Total	5.45	4	0.2446

62. dwstat

Durbin-Watson d-statistic(2, 11) = 1.381133

63. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	0.008	1	0.9275

H0: no serial correlation

64. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	0.113	2	0.9449

H0: no serial correlation

65. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn
 Ho: model has no omitted variables
 F(3, 6) = 1.52
 Prob > F = 0.3018

66. estat vif

Variable	VIF	1/VIF
g	1.00	1.000000
Mean VIF	1.00	

67. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-109.1314	2	222.2628	223.0586

Note: N=Obs used in calculating BIC; see **[R] BIC note**

68.

69. reg jgn h

Source	SS	df	MS			
Model	1.2376e+09	1	1.2376e+09	Number of obs =	11	
Residual	497631674	9	55292408.3	F(1, 9) =	22.38	
Total	1.7353e+09	10	173525941	Prob > F =	0.0011	
				R-squared =	0.7132	
				Adj R-squared =	0.6814	
				Root MSE =	7435.9	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
h	.9056124	.1914168	4.73	0.001	.4725975	1.338627
_cons	15141.06	5743.309	2.64	0.027	2148.787	28133.32

70. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance
Variables: fitted values of jgn

chi2(1) = 1.43
Prob > chi2 = 0.2319

71. estat imtest, white

White's test for Ho: homoskedasticity
against Ha: unrestricted heteroskedasticity

chi2(2) = 1.79
Prob > chi2 = 0.4091

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	1.79	2	0.4091
Skewness	2.44	1	0.1180
Kurtosis	2.70	1	0.1003
Total	6.93	4	0.1395

72. dwstat

Durbin-Watson d-statistic(2, 11) = .8948624

73. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	2.202	1	0.1378

H0: no serial correlation

74. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	2.206	2	0.3318

H0: no serial correlation

75. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn

Ho: model has no omitted variables

F(3, 6) = 0.67
 Prob > F = 0.6014

76. estat vif

Variable	VIF	1/VIF
h	1.00	1.000000
Mean VIF	1.00	

77. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-112.5594	2	229.1189	229.9147

Note: N=Obs used in calculating BIC; see [R] BIC note

78.

79. reg jgn total

Source	SS	df	MS	Number of obs = 11		
Model	1.3346e+09	1	1.3346e+09	F(1, 9) =	29.98	
Residual	400698283	9	44522031.5	Prob > F =	0.0004	
Total	1.7353e+09	10	173525941	R-squared =	0.7691	
				Adj R-squared =	0.7434	
				Root MSE =	6672.5	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
total	.6266005	.1144482	5.47	0.000	.3677007	.8855003
_cons	8823.294	6066.456	1.45	0.180	-4899.983	22546.57

80. estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
 Ho: Constant variance
 Variables: fitted values of jgn

 chi2(1) = 1.57
 Prob > chi2 = 0.2097

81. estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

 chi2(2) = 1.43
 Prob > chi2 = 0.4885

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	1.43	2	0.4885
Skewness	2.97	1	0.0849
Kurtosis	2.61	1	0.1064
Total	7.01	4	0.1355

82. dwstat

Durbin-Watson d-statistic(2, 11) = 1.007931

83. estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	1.138	1	0.2860

H0: no serial correlation

84. estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	1.171	2	0.5568

H0: no serial correlation

85. estat ovtest

Ramsey RESET test using powers of the fitted values of jgn
 Ho: model has no omitted variables
 F(3, 6) = 0.59
 Prob > F = 0.6443

86. estat vif

Variable	VIF	1/VIF
total	1.00	1.000000
Mean VIF	1.00	

87. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	11	-119.4292	-111.3679	2	226.7357	227.5315

Note: N=Obs used in calculating BIC; see **[R] BIC note**

88.

89. *no autocoreelation or heterskedasticity detected causing false possitive relationsh
> ips in those HIA categories providing signifcant predictions of JGN dwellings

90.

91. reg jgn L1.a

Source	SS	df	MS	Number of obs = 10		
Model	1.5113e+09	1	1.5113e+09	F(1, 8) =	234.23	
Residual	51619469	8	6452433.63	Prob > F =	0.0000	
Total	1.5630e+09	9	173662406	R-squared =	0.9670	
				Adj R-squared =	0.9628	
				Root MSE =	2540.2	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
a						
L1.	2.078286	.1357956	15.30	0.000	1.765141	2.391431
_cons	-3667.671	3052.884	-1.20	0.264	-10707.64	3372.293

92. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-91.47351	2	186.947	187.5522

Note: N=Obs used in calculating BIC; see **[R] BIC note**

93. reg jgn L1.b

Source	SS	df	MS	Number of obs = 10		
Model	695765079	1	695765079	F(1, 8) =	6.42	
Residual	867196574	8	108399572	Prob > F =	0.0351	
Total	1.5630e+09	9	173662406	R-squared =	0.4452	
				Adj R-squared =	0.3758	
				Root MSE =	10412	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
b						
L1.	13.03288	5.14426	2.53	0.035	1.170196	24.89556
_cons	8370.548	13449.9	0.62	0.551	-22644.97	39386.07

94. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-105.5803	2	215.1607	215.7659

Note: N=Obs used in calculating BIC; see **[R] BIC note**

95. reg jgn L1.c

Source	SS	df	MS	Number of obs = 10		
Model	1.4953e+09	1	1.4953e+09	F(1, 8) =	176.79	
Residual	67663087.5	8	8457885.93	Prob > F =	0.0000	
Total	1.5630e+09	9	173662406	R-squared =	0.9567	
				Adj R-squared =	0.9513	
				Root MSE =	2908.2	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
^c L1.	8.762547	.6590181	13.30	0.000	7.242849	10.28225
_cons	9554.412	2566.185	3.72	0.006	3636.78	15472.04

96. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-92.82672	2	189.6534	190.2586

Note: N=Obs used in calculating BIC; see **[R] BIC note**

97. reg jgn L1.d

Source	SS	df	MS	Number of obs = 10		
Model	102034323	1	102034323	F(1, 8) =	0.56	
Residual	1.4609e+09	8	182615916	Prob > F =	0.4762	
Total	1.5630e+09	9	173662406	R-squared =	0.0653	
				Adj R-squared =	-0.0516	
				Root MSE =	13514	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
^d L1.	-3.863883	5.169161	-0.75	0.476	-15.78399	8.056225
_cons	46960.53	8568.703	5.48	0.001	27201.06	66719.99

98. estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-108.1881	2	220.3763	220.9815

Note: N=Obs used in calculating BIC; see **[R] BIC note**

99. reg jgn L1.e

Source	SS	df	MS			
Model	879747393	1	879747393	Number of obs =	10	
Residual	683214259	8	85401782.4	F(1, 8) =	10.30	
Total	1.5630e+09	9	173662406	Prob > F =	0.0124	
				R-squared =	0.5629	
				Adj R-squared =	0.5082	
				Root MSE =	9241.3	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
L1.e	34.58198	10.77468	3.21	0.012	9.735526	59.42843
_cons	4395.808	11896.65	0.37	0.721	-23037.93	31829.54

100 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-104.3881	2	212.7761	213.3813

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

101 reg jgn L1.f

Source	SS	df	MS			
Model	1.4201e+09	1	1.4201e+09	Number of obs =	10	
Residual	142829485	8	17853685.6	F(1, 8) =	79.54	
Total	1.5630e+09	9	173662406	Prob > F =	0.0000	
				R-squared =	0.9086	
				Adj R-squared =	0.8972	
				Root MSE =	4225.4	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
L1.f	1.175877	.1318443	8.92	0.000	.8718434	1.479911
_cons	21007.32	2649.166	7.93	0.000	14898.33	27116.31

102 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-96.56227	2	197.1245	197.7297

Note: N=Obs used in calculating BIC; see [\[R\] BIC note](#)

103 reg jgn L1.g

Source	SS	df	MS			
Model	1.5125e+09	1	1.5125e+09	Number of obs =	10	
Residual	50466139.5	8	6308267.44	F(1, 8) =	239.76	
Total	1.5630e+09	9	173662406	Prob > F =	0.0000	
				R-squared =	0.9677	
				Adj R-squared =	0.9637	
				Root MSE =	2511.6	

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
jgn						
g						
L1.	2.081689	.1344386	15.48	0.000	1.771673	2.391705
_cons	-3814.964	3026.693	-1.26	0.243	-10794.53	3164.603

104 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-91.36053	2	186.7211	187.3262

Note: N=Obs used in calculating BIC; see **[R] BIC note**

105 reg jgn L1.h

Source	SS	df	MS	Number of obs = 10	
Model	1.4601e+09	1	1.4601e+09	F(1, 8) =	113.52
Residual	102895128	8	12861890.9	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9342
				Adj R-squared =	0.9259
				Root MSE =	3586.3

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
jgn						
h						
L1.	.9990346	.0937663	10.65	0.000	.7828092	1.21526
_cons	14459.14	2772.031	5.22	0.001	8066.827	20851.46

106 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-94.92256	2	193.8451	194.4503

Note: N=Obs used in calculating BIC; see **[R] BIC note**

107 reg jgn L1.total

Source	SS	df	MS	Number of obs = 10	
Model	1.4908e+09	1	1.4908e+09	F(1, 8) =	165.26
Residual	72165738.6	8	9020717.32	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9538
				Adj R-squared =	0.9481
				Root MSE =	3003.5

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
jgn						
total						
L1.	.6813358	.0529996	12.86	0.000	.5591185	.8035531
_cons	8227.573	2750.302	2.99	0.017	1885.366	14569.78

108 estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
 Ho: Constant variance
 Variables: fitted values of jgn

 chi2(1) = **4.74**
 Prob > chi2 = **0.0295**

109 estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

 chi2(2) = **3.65**
 Prob > chi2 = **0.1611**

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	3.65	2	0.1611
Skewness	6.06	1	0.0138
Kurtosis	2.65	1	0.1036
Total	12.36	4	0.0149

110 dwstat

Durbin-Watson d-statistic(2, 10) = **1.685001**

111 estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	1.401	1	0.2366

H0: no serial correlation

112 estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	2.471	2	0.2906

H0: no serial correlation

113 estat ovtest

Ramsey RESET test using powers of the fitted values of jgn
 Ho: model has no omitted variables
 F(3, 5) = **0.18**
 Prob > F = **0.9089**

114 estat vif

Variable	VIF	1/VIF
total L1.	1.00	1.000000
Mean VIF	1.00	

115 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-93.14884	2	190.2977	190.9028

Note: N=Obs used in calculating BIC; see **[R] BIC note**

116 *All 1 year lags provide significant forecasts of JGN commencements (other than D)
 117 *note that range of AIC value is small, but A,G,H providing the best comparative power
 > er
 118 *AIC is lower than in all unlagged regressions
 119
 120 reg jgn a L1.a

Source	SS	df	MS	Number of obs =	10
Model	1.5114e+09	2	755678965	F(2, 7) =	102.51
Residual	51603723.5	7	7371960.5	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9670
				Adj R-squared =	0.9576
				Root MSE =	2715.1

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
a					
--.	-0.0206535	.4468967	-0.05	0.964	-1.077396 1.036089
L1.	2.096582	.4216519	4.97	0.002	1.099534 3.09363
_cons	-3584.61	3725.377	-0.96	0.368	-12393.73 5224.507

121 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-91.47198	3	188.944	189.8517

Note: N=Obs used in calculating BIC; see **[R] BIC note**

122 reg jgn b L1.b

Source	SS	df	MS	Number of obs =	10
Model	924625611	2	462312805	F(2, 7) =	5.07
Residual	638336042	7	91190863.1	Prob > F =	0.0435
Total	1.5630e+09	9	173662406	R-squared =	0.5916
				Adj R-squared =	0.4749
				Root MSE =	9549.4

	jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
b							
--.		8.524357	5.380862	1.58	0.157	-4.19936	21.24807
L1.		12.42297	4.733972	2.62	0.034	1.228909	23.61704
_cons		-12537.35	18065.51	-0.69	0.510	-55255.49	30180.79

123 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-104.0483	3	214.0967	215.0044

Note: N=Obs used in calculating BIC; see **[R] BIC note**

124 reg jgn c L1.c

Source	SS	df	MS	Number of obs = 10		
Model	1.5259e+09	2	762943343	F(2, 7) =	144.05	
Residual	37074966.2	7	5296423.75	Prob > F =	0.0000	
Total	1.5630e+09	9	173662406	R-squared =	0.9763	
				Adj R-squared =	0.9695	
				Root MSE =	2301.4	

	jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
c							
--.		3.749656	1.560294	2.40	0.047	.0601474	7.439165
L1.		4.763389	1.743917	2.73	0.029	.6396804	8.887099
_cons		8824.7	2053.288	4.30	0.004	3969.445	13679.95

125 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-89.81872	3	185.6374	186.5452

Note: N=Obs used in calculating BIC; see **[R] BIC note**

126 reg jgn d L1.d

Source	SS	df	MS	Number of obs = 10		
Model	809119140	2	404559570	F(2, 7) =	3.76	
Residual	753842513	7	107691788	Prob > F =	0.0779	
Total	1.5630e+09	9	173662406	R-squared =	0.5177	
				Adj R-squared =	0.3799	
				Root MSE =	10377	

	jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
d							
--.		-10.49065	4.094097	-2.56	0.037	-20.17166	-.8096536
L1.		-5.89042	4.047576	-1.46	0.189	-15.46142	3.680577
_cons		65057.48	9652.881	6.74	0.000	42232.04	87882.91

127 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-104.8799	3	215.7599	216.6676

Note: N=Obs used in calculating BIC; see **[R] BIC note**

128 reg jgn e Ll.e

Source	SS	df	MS	Number of obs = 10		
Model	922461923	2	461230962	F(2, 7) =	5.04	
Residual	640499730	7	91499961.4	Prob > F =	0.0441	
Total	1.5630e+09	9	173662406	R-squared =	0.5902	
				Adj R-squared =	0.4731	
				Root MSE =	9565.6	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
e						
--.	8.340249	12.2068	0.68	0.516	-20.52424	37.20473
L1.	30.72207	12.50195	2.46	0.044	1.159642	60.28449
_cons	-139.2802	13989.06	-0.01	0.992	-33218.14	32939.58

129 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-104.0653	3	214.1305	215.0383

Note: N=Obs used in calculating BIC; see **[R] BIC note**

130 reg jgn f Ll.f

Source	SS	df	MS	Number of obs = 10		
Model	1.4357e+09	2	717871485	F(2, 7) =	39.50	
Residual	127218684	7	18174097.7	Prob > F =	0.0002	
Total	1.5630e+09	9	173662406	R-squared =	0.9186	
				Adj R-squared =	0.8953	
				Root MSE =	4263.1	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
f						
--.	-.3268195	.352632	-0.93	0.385	-1.160662	.5070228
L1.	1.448472	.322807	4.49	0.003	.6851548	2.211789
_cons	22547.8	3147.502	7.16	0.000	15105.14	29990.46

131 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-95.98355	3	197.9671	198.8749

Note: N=Obs used in calculating BIC; see **[R] BIC note**

132 reg jgn g Ll.g

Source	SS	df	MS	Number of obs =	10
Model	1.5125e+09	2	756253335	F(2, 7) =	104.92
Residual	50454983	7	7207854.71	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9677
				Adj R-squared =	0.9585
				Root MSE =	2684.7

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
g					
--.	-0.0171794	.4366634	-0.04	0.970	-1.049724 1.015365
L1.	2.096862	.4115575	5.09	0.001	1.123683 3.070041
_cons	-3744.812	3694.156	-1.01	0.344	-12480.1 4990.478

133 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-91.35942	3	188.7188	189.6266

Note: N=Obs used in calculating BIC; see **[R] BIC note**

134 reg jgn h Ll.h

Source	SS	df	MS	Number of obs =	10
Model	1.4707e+09	2	735368116	F(2, 7) =	55.82
Residual	92225419.9	7	13175060	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9410
				Adj R-squared =	0.9241
				Root MSE =	3629.7

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
h					
--.	-0.214174	.2379944	-0.90	0.398	-0.7769413 0.3485933
L1.	1.172176	.2145307	5.46	0.001	0.6648919 1.679461
_cons	16070.29	3328.148	4.83	0.002	8200.47 23940.11

135 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-94.37519	3	194.7504	195.6581

Note: N=Obs used in calculating BIC; see **[R] BIC note**

136 reg jgn total L1.total

Source	SS	df	MS	Number of obs = 10		
Model	1.4991e+09	2	749573725	F(2, 7) =	82.22	
Residual	63814203.9	7	9116314.84	Prob > F =	0.0000	
Total	1.5630e+09	9	173662406	R-squared =	0.9592	
				Adj R-squared =	0.9475	
				Root MSE =	3019.3	

jgn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
total						
---	-.1428118	.1492075	-0.96	0.370	-.4956314	.2100078
L1.	.8008558	.135764	5.90	0.001	.4798248	1.121887
_cons	9918.89	3281.284	3.02	0.019	2159.886	17677.89

137 estat hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance
 Variables: fitted values of jgn

chi2(1) = 1.76
 Prob > chi2 = 0.1842

138 estat imtest, white

White's test for Ho: homoskedasticity
 against Ha: unrestricted heteroskedasticity

chi2(5) = 9.39
 Prob > chi2 = 0.0944

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	9.39	5	0.0944
Skewness	8.93	2	0.0115
Kurtosis	0.09	1	0.7643
Total	18.42	8	0.0183

139 dwstat

Durbin-Watson d-statistic(3, 10) = 2.083126

140 estat bgodfrey, lags(1)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
1	1.899	1	0.1682

H0: no serial correlation

141 estat bgodfrey, lags(2)

Breusch-Godfrey LM test for autocorrelation

lags(p)	chi2	df	Prob > chi2
2	3.526	2	0.1715

H0: no serial correlation

142 estat ovtest

Ramsey RESET test using powers of the fitted values of jgn

Ho: model has no omitted variables

F(3, 4) = **0.21**
 Prob > F = **0.8836**

143 estat vif

Variable	VIF	1/VIF
total		
L1.	6.49	0.154012
--.	6.49	0.154012
Mean VIF	6.49	

144 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-92.53389	3	191.0678	191.9755

Note: N=Obs used in calculating BIC; see **[R] BIC note**

145

146 *All 1 year lags provide significant forecasts

147 *Unlagged HIA categories fail to provide significant relationships when combined with > h lags, P Value exceeds acceptable threshold

148 *Unlagged categories should therefore be dropped from the model.

149

150 *final two candidates for models providing the highest comparative power

151 reg jgn L1.a

Source	SS	df	MS	Number of obs =	10
Model	1.5113e+09	1	1.5113e+09	F(1, 8) =	234.23
Residual	51619469	8	6452433.63	Prob > F =	0.0000
				R-squared =	0.9670
				Adj R-squared =	0.9628
Total	1.5630e+09	9	173662406	Root MSE =	2540.2

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
jgn						
a						
L1.	2.078286	.1357956	15.30	0.000	1.765141	2.391431
_cons	-3667.671	3052.884	-1.20	0.264	-10707.64	3372.293

152 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-91.47351	2	186.947	187.5522

Note: N=Obs used in calculating BIC; see **[R] BIC note**

153

154 reg jgn L1.total

Source	SS	df	MS	Number of obs =	10
Model	1.4908e+09	1	1.4908e+09	F(1, 8) =	165.26
Residual	72165738.6	8	9020717.32	Prob > F =	0.0000
Total	1.5630e+09	9	173662406	R-squared =	0.9538
				Adj R-squared =	0.9481
				Root MSE =	3003.5

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
jgn						
total						
L1.	.6813358	.0529996	12.86	0.000	.5591185	.8035531
_cons	8227.573	2750.302	2.99	0.017	1885.366	14569.78

155 estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	10	-108.5257	-93.14884	2	190.2977	190.9028

Note: N=Obs used in calculating BIC; see **[R] BIC note**

156

157 log close

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log type: smcl
closed on: 10 Dec 2019, 09:32:21